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CHOOSING A CONSTANT ‘ $\alpha$ ’ IN EXPONENTIAL SMOOTHING USING ARMA MODEL<br>P. Ramakrishna Reddy, B. Sarojamma*<br>Department of Statistics, S.V. University, Tirupati - 517502, Andhra Pradesh, India


#### Abstract

Time series analysis and forecasting process plays an important role in business, atmospheric studies, insurance companies, banking sectors etc. There are many forecasting techniques like moving averages, double moving average, multiple moving averages, simple exponential smoothing, adaptive smoothing, double exponential smoothing, triple exponential smoothing, autoregressive integrated moving averages, etc. In simple exponential smoothing, constant ' $\alpha$ ' is not fixed, it may vary from 0 to 1 . In this paper, we discuss about ' $\alpha$ ' where it is estimated through some process. We estimating the constant of exponential smoothing using autoregressive moving average models by various autoregressive and moving average parameters. We fitting a new exponential smoothing model by fixing value to ' $\alpha$ '. To check for a goodness of fit, we use Kolmogrov - Smirnov test to simple exponential smoothing and new exponential smoothing models. Mean square error criteria are used for the purpose of choosing best model between simple exponential smoothing model and new exponential smoothing model.


KEY WORDS: Simple Exponential Smoothing, Autoregressive Moving Average, Errors, Absolute errors, Mean Square Error.

## INTRODUCTION

Forecasting plays an important role in national economy, customers, products in industry, marketing organization, executive offices, atmospheric studies, financial organizations, etc. There are many forecasting models like moving averages, exponential smoothing, auto regression, autoregressive moving averages (ARMA), autoregressive
integrated moving averages (ARIMA), vector autoregressive models, autoregressive conditional heteroscedasticity (ARCH), generalized autoregressive conditional heteroscedasticity (GARCH), etc. In moving averages, we take same weight for each observation.

Exponential smoothing is a procedure for continually revising a forecast in the light of more recent emporia. Exponential smoothing assigns exponentially decreasing weights as the observation get older. In other words, relatively more weights are given for recent observations in forecasting than the older observations. Exponential smoothing is a simple technique used to smooth and forecast a time-series without the necessity of fitting a parametric model.

## SIMPLE EXPONENTIAL SMOOTHING

Simple exponential smoothing is also called a single exponential smoothing. The parameter in simple exponential smoothing is ' $\alpha$ '. Simple exponential smoothing is used for short range forecasting, usually just for one month to the future. The model of the simple exponential smoothing is

$$
S_{t}=Y_{t}+(1-\alpha) S_{t-1}
$$

where $S_{t}=$ forecast for time ' $t$ '
$\mathrm{S}_{\mathrm{t}-1}=$ forecast for time ' $\mathrm{t}-1$,
$\mathrm{Y}_{\mathrm{t}}=$ time series observation at time ' $t$ '

$$
\alpha=\text { constant }
$$

$\alpha$ lies between 0 and1
$\alpha+\beta=1$

DOUBLE EXPONENTIAL SMOOTHING

It is an extension of single exponential smoothing method. If we compute single exponential smoothing to earlier single exponential smoothing, then we get double exponential smoothing. Under double exponential smoothing, the two smoothing averages are defined by

$$
\begin{aligned}
& \mathrm{S}_{\mathrm{T}}=\alpha \mathrm{Y}_{\mathrm{T}}+(1 \\
& -\alpha) \mathrm{S}_{\mathrm{T}-1} \\
& \mathrm{~S}_{\mathrm{T}}^{(2)}=\alpha \mathrm{S}_{\mathrm{T}}+ \\
& (1-\alpha) \mathrm{S}_{\mathrm{T}-1}^{(2)} \\
& \text { where } \quad \mathrm{T} \quad \text { is } \\
& \text { current } \quad \text { time } \\
& \text { period }
\end{aligned}
$$

## TRIPLE EXPONENTIAL SMOOTHING

It is an extension to double exponential smoothing. If we compute single exponential Smoothing to double exponential smoothing, then we get triple exponential smoothing, the three smoothing averages are defined by

$$
\begin{aligned}
& \mathrm{S}_{\mathrm{T}}=\alpha \mathrm{Y}_{\mathrm{T}}+(1 \\
& -\alpha) \mathrm{S}_{\mathrm{T}-1} \\
& \mathrm{~S}_{\mathrm{T}}^{(2)}=\alpha \mathrm{S}_{\mathrm{T}}+ \\
& (1-\alpha) \mathrm{S}_{\mathrm{T}-1}^{(2)} \\
& \mathrm{S}_{\mathrm{T}}^{(3)}=\alpha \mathrm{S}_{\mathrm{T}}^{(2)}+ \\
& (1-\alpha) \mathrm{S}_{\mathrm{T}-1}^{(3)}
\end{aligned}
$$

## ARMA (AUTO REGRESSIVE MOVING AVERAGE)

The combination of auto regression of ' $p$ ' recent time-series observations $\mathrm{Y}_{\mathrm{t}-1}, \mathrm{Y}_{\mathrm{t}-2} \ldots \mathrm{Y}_{\mathrm{t}-\mathrm{p}}$ with moving averages of a recent error terms $e_{t}, e_{t-1} \ldots e_{t-q}$ is known as ARMA (p,q) The auto regressive moving average model of order $(\mathrm{p}, \mathrm{q})$ is $Y_{t}=c+e_{t}-\theta_{1} e_{t-1}-\theta_{2} e_{t-2}-\cdots-\theta_{q} e_{t-q}+\phi_{1} Y_{t-1}+\phi_{2} Y_{t-2}+\cdots+\phi_{p} Y_{t-p}$ In this paper, we estimated the parameter of simple exponential smoothing by using ARMA models. After estimating parameter, we fitted a new exponential smoothing and comparing with simple exponential smoothing based on forecasting accuracy mean square error (MSE)

## METHODOLOGY

By using time series observations, we forecast future values. There are many models for forecasting, we now fitting an equation by using ARMA. In general, the time series data is in the form

| t | $\mathrm{t}_{1}$ | $\mathrm{t}_{2}$ | $\mathrm{t}_{3}$ | $\ldots$ | $\mathrm{t}_{\mathrm{n}}$ |
| :--- | :--- | :--- | :--- | :--- | :--- |
| $\mathrm{Y}_{\mathrm{t}}$ |  | $\mathrm{Y}_{1}$ | $\mathrm{Y}_{2}$ | $\mathrm{Y}_{3}$ | $\ldots$ |
| $\mathrm{Y}_{\mathrm{n}}$ |  |  |  |  |  |

## AUTO REGRESSION (AR)

$\mathrm{p}^{\text {th }}$ order auto regression equation is formed by taking regression equation of time series values of order $t-1, t-2, \ldots$ $\mathrm{t}-\mathrm{p}$ with their $1,2, \ldots \mathrm{p}^{\text {th }}$ order auto regressive parameter, added with a constant term ' $c$ ' and also error term ' $e_{t}$ '.
$p^{\text {th }}$ order AR model is

$$
\mathrm{Y}_{\mathrm{t}}=\mathrm{c}+\phi_{1} \mathrm{Y}_{\mathrm{t}-1}+\phi_{2} \mathrm{Y}_{\mathrm{t}-2}+\cdots+\phi_{\mathrm{p}} \mathrm{Y}_{\mathrm{t}-\mathrm{p}}
$$

where $\mathrm{c}=$ constant term,
$\phi_{p}=p^{\text {th }}$ autoregressive parameter

$$
e_{t}=\text { the error term at }
$$

time ' $t$ '.

## MOVING AVERAGE

Moving average is the relation of error terms with time series value. The $\mathrm{q}^{\text {th }}$ order moving average model is multiplication of error terms of time points $\mathrm{t}-1, \mathrm{t}-2, \ldots \mathrm{t}-\mathrm{q}$ with $\theta_{1}, \theta_{2}, \ldots, \theta_{\mathrm{q}}$ moving average parameters subtracted from error term ' $e_{t}$ ' and constant ' $c$ '.
The $q^{\text {th }}$ order moving average model is

$$
\mathrm{Y}_{\mathrm{t}}=c+\mathrm{e}_{\mathrm{t}}-\theta_{1} \mathrm{e}_{\mathrm{t}-1}-\theta_{2} \mathrm{e}_{\mathrm{t}-2}-\cdots-\theta_{\mathrm{q}} \mathrm{e}_{\mathrm{t}-\mathrm{q}}
$$ where

$c=$ constant term
$\theta_{q}=q^{\text {th }}$ moving average parameter
$e_{t-k}=$ the error term at time $t-k$

## AUTO REGRESSIVE MOVING AVERAGE MODEL (p, q)

Auto regressive moving average model ( $p, q$ ) is a mixture of auto regression of order ' $p$ ' and moving averages of order ' $q$ '. The equation of ARMA $(p, q)$ is
$\mathrm{Y}_{\mathrm{t}}=c+\phi_{1} \mathrm{Y}_{\mathrm{t}-1}+\cdots+\phi_{\mathrm{p}} \mathrm{Y}_{\mathrm{t}-\mathrm{p}}+\mathrm{e}_{\mathrm{t}}-\theta_{1} \mathrm{e}_{\mathrm{t}-1}-\cdots-\theta_{\mathrm{q}} \mathrm{e}_{\mathrm{t}-\mathrm{q}}$
where
models. A model possessing minimum MSE value that
$\mathrm{Y}_{\mathrm{t}}=$ Forecast time series value at time ' t '
$\phi_{1}, \phi_{2}, \cdots \phi_{\mathrm{p}}=$ Auto regressive parameters
model is the best model among these fitted nine models.
$\theta_{1}, \theta_{2}, \cdots \theta_{\mathrm{q}}=$ Moving average parameters

$e_{t-1}, e_{t-2}, \cdots e_{t-q}$ are error terms at time points $(t-1)^{t h},(t-2)^{t h} \cdots(t-q)^{\text {th }}$.
$B, B^{2}, \cdots B^{p}$ are back shift operators of time series values
$B, B^{2}, \cdots B^{q}$ are back shift operators of error terms.
$\operatorname{MSE}=\frac{\sum_{i=1}^{\mathrm{n}}\left(\mathrm{Y}_{\mathrm{t}}-\hat{\mathrm{Y}}_{\mathrm{t}}\right)^{2}}{\mathrm{n}}$
where
$Y_{t}$ : actual value of time point ' t '
$\hat{\mathrm{Y}}_{\mathrm{t}}$ : estimated value of time point $\mathrm{t}^{\prime}$ n : number of time points

By calculating the error and absolute errors of the model, we fitting a new exponential smoothing model.
Generally a simple exponential smoothing equation is given as a forecast at time point ${ }^{〔} t+1$ ' and is a combination of time series value at time point $t$ is $Y_{t}$ and forecast at time point $t$ i.e., $\mathrm{F}_{\mathrm{t}}$ with constant terms $\alpha$ and $\beta$.

$$
\mathrm{F}_{\mathrm{t}+1}=\alpha \mathrm{Y}_{\mathrm{t}}+(1-\alpha) \mathrm{F}_{\mathrm{t}}
$$

where $\alpha$ is a smoothing constant lies between 0 and 1

$$
\begin{aligned}
\beta & =1-\alpha \\
Y_{t} & =\text { time series }
\end{aligned}
$$ value at time ' $t$ '

$$
F_{t}=\text { forecast }
$$ values at time' $t$ '

$$
F_{t+1}=\text { forecast }
$$ values at time $\mathrm{t}+1$

$\mathrm{Y}_{\mathrm{t}}=\mathrm{c}+\phi_{1} \mathrm{Y}_{\mathrm{t}-1}+\phi_{2} \mathrm{Y}_{\mathrm{t}-2}+\phi_{3} \mathrm{Y}_{\mathrm{t}-3}+\mathrm{e}_{\mathrm{t}}-\theta_{1} \mathrm{e}_{\mathrm{t}-1}-\theta_{2} \mathrm{e}_{\mathrm{t}-2}-\theta_{3} \mathrm{e}_{\mathrm{t}-3}$
In simple exponential smoothing $\alpha$ is not a fixed value, it may various between 0 and 1 . In our new exponential smoothing model, we fitting the parameter ' $\alpha$ ' by using autoregressive moving average models. For the purpose of fixing ' $\alpha$ ' we fitted nine ARMA models i.e. ARMA ( 1,1 ), ARMA (1, 2), ARMA (1, 3), ARMA (2, 1), ARMA (2, 2), ARMA $(2,3)$, ARMA $(3,1)$, ARMA (3, 2), ARMA (3, 3). We computed mean square error values for the nine ARMA models, from that we estimating ' $\alpha$ "' value.
We fitting new exponential smoothing equation of the form
$\mathrm{F}_{\mathrm{t}+1}{ }^{*}=\alpha^{*} \mathrm{Y}_{\mathrm{t}}+\left(1-\alpha^{*}\right) \mathrm{F}_{\mathrm{t}}$
where
$\alpha^{*}$ is estimated by using the formulae
$\alpha^{*}=\frac{\sum \varepsilon_{\mathrm{t}}}{\sum\left|\varepsilon_{\mathrm{t}}\right|}$
$\sum\left|\varepsilon_{\mathrm{t}}\right|$ is sum of absolute errors of $\operatorname{ARMA}(3,2)$ model $\sum \varepsilon_{\mathrm{t}}$ is sum of errors of $\operatorname{ARMA}(3,2)$ model

The sum of squares of errors divided by number of observations gives MSE. MSE is calculated for two models i.e. exponential smoothing model and new exponential smoothing model.
MSE of simple exponential smoothing model is

$$
\operatorname{MSESES}=\frac{\sum_{t=1}^{n}\left(Y_{t}-F_{t}\right)^{2}}{n}
$$

MSE of new exponential smoothing model is

$$
\text { MSENES }=\frac{\sum_{t=1}^{n}\left(Y_{t}-F_{t}^{*}\right)^{2}}{n}
$$

## EMPIRICAL INVESTIGATIONS

In single exponential smoothing model, $\alpha$ is not a fixed value. It may take value from 0 to 1 . By using ARMA models, we estimated $\alpha^{*}$ value, and we fitting new exponential smoothing model by using estimated ' $\alpha$ ',
We fitted nine ARMA models for the given data. Various models are
ARMA $(1,1)$ Model: $\quad Y_{t}+0.997 Y_{t-1}-\varepsilon_{t}-0.978 \varepsilon_{t-1}+$ $122.722=0$

ARMA $(1,2)$ Model: $\quad Y_{t}-0.460 Y_{t-1}-\varepsilon_{t}+0.677 \varepsilon_{t-1}+$ $0.321 \varepsilon_{\mathrm{t}-2}+120.221=0$

| Annual | ARMA <br> $(\mathbf{1 , 1})$ | ARMA <br> $(\mathbf{1 , 2})$ | ARMA <br> $(\mathbf{1 , 3})$ | ARMA <br> $\mathbf{( 2 , 1 )}$ | ARMA <br> $(\mathbf{2 , 2})$ | ARMA <br> $(\mathbf{2 , 3})$ | ARMA <br> $(\mathbf{3 , 1})$ | ARMA <br> $(\mathbf{3 , 2})$ | ARMA <br> $(\mathbf{3 , 3})$ |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 64.4 | 0.0217 | 0.0099 | 0.0054 | 0.0071 | 0.0161 | 0.0048 | 0.0155 | 0.0072 | 0.0112 |
| 66.2 | 2.3963 | 2.5773 | 2.6798 | 2.6549 | 2.5233 | 2.6945 | 2.5335 | 2.6436 | 2.5796 |
| 64.4 | 0.0531 | 0.0768 | 0.1665 | 0.2338 | 0.2445 | 0.1981 | 0.2506 | 0.1659 | 0.1902 |
| 63.9 | 1.1057 | 0.2362 | 0.3525 | 0.1961 | 0.3502 | 0.2309 | 0.4191 | 0.4554 | 0.4740 |
| 66.8 | 3.7415 | 3.9780 | 6.4034 | 4.8272 | 4.7873 | 6.3142 | 5.3870 | 6.4277 | 6.1385 |
| 65.5 | 0.2706 | 0.3868 | 0.0241 | 0.2453 | 0.3949 | 0.0179 | 0.3657 | 0.0968 | 0.0893 |
| 65.4 | 0.0722 | 1.0248 | 0.4889 | 1.0329 | 0.6854 | 0.6889 | 0.4050 | 0.7702 | 0.5640 |
| 63.3 | 3.5446 | 1.5488 | 0.3009 | 1.0885 | 1.4040 | 0.1960 | 1.1270 | 0.9964 | 1.3017 |


| 65.7 | 0.0783 | 0.7508 | 0.3901 | 1.3832 | 0.8131 | 0.4811 | 0.8940 | 1.1124 | 0.8571 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 65.1 | 0.0303 | 0.0883 | 0.0001 | 0.0401 | 0.1633 | 0.0090 | 0.0838 | 0.0136 | 0.0871 |
| 66.2 | 0.3266 | 0.9094 | 0.4153 | 0.8979 | 0.1830 | 0.5483 | 0.1058 | 0.2397 | 0.0965 |
| 65.5 | 0.0044 | 0.0124 | 0.1347 | 0.0175 | 0.1261 | 0.1355 | 0.0376 | 0.0561 | 0.0015 |
| 65.7 | 0.0190 | 0.1979 | 0.0745 | 0.2501 | 0.0001 | 0.1071 | 0.0047 | 0.0000 | 0.0139 |
| 64.5 | 1.2819 | 0.8896 | 0.5491 | 0.6706 | 1.4216 | 0.4761 | 1.1185 | 0.4348 | 0.6574 |
| 66.8 | 0.5435 | 1.2102 | 1.4542 | 1.5906 | 0.8740 | 1.5124 | 0.9147 | 0.7437 | 0.6419 |
| 66.1 | 0.1192 | 0.0201 | 0.0381 | 0.0195 | 0.0062 | 0.0250 | 0.0011 | 0.0561 | 0.0155 |
| 65.9 | 0.1335 | 0.0562 | 0.0136 | 0.0519 | 0.0323 | 0.0013 | 0.1014 | 0.1648 | 0.2507 |
| 65.4 | 0.3250 | 0.2048 | 0.0022 | 0.1025 | 0.4265 | 0.0080 | 0.3196 | 0.0088 | 0.0447 |
| 66.1 | 0.1296 | 0.0005 | 0.0032 | 0.0234 | 0.0133 | 0.0004 | 0.0058 | 0.1248 | 0.1408 |
| 66 | 0.0236 | 0.1316 | 0.1082 | 0.0912 | 0.2688 | 0.1386 | 0.2506 | 0.0072 | 0.0269 |
| 65.8 | 0.7080 | 0.4436 | 0.3355 | 0.4290 | 0.8328 | 0.3140 | 0.9128 | 1.3577 | 1.3884 |
| 64.5 | 3.4808 | 5.0024 | 4.5950 | 4.8960 | 5.7730 | 4.6423 | 5.5937 | 3.5261 | 3.6355 |
| 68.3 | 2.0739 | 1.0268 | 1.4709 | 1.0968 | 1.0681 | 1.2199 | 1.1809 | 0.8690 | 1.1002 |
| 68.6 | 4.5882 | 1.2224 | 1.1402 | 0.6803 | 1.0217 | 0.8870 | 1.1462 | 1.1903 | 1.4643 |
| 67.4 | 0.1382 | 0.6125 | 0.0003 | 0.2107 | 0.1741 | 0.0026 | 0.0367 | 0.0218 | 0.0053 |
| 65.6 | 1.2113 | 0.4801 | 0.0227 | 0.4160 | 0.3904 | 0.0023 | 0.3446 | 0.0777 | 0.0644 |
| 65.3 | 3.8322 | 1.5987 | 1.2701 | 0.9930 | 0.6410 | 1.0276 | 0.5510 | 1.5770 | 1.3661 |
| 66 | 0.8365 | 1.7020 | 1.5036 | 1.1550 | 0.6129 | 1.7361 | 0.6501 | 0.2356 | 0.2281 |
| 66.5 | 0.8403 | 1.9113 | 1.3811 | 1.9179 | 1.4417 | 1.6266 | 1.7348 | 2.5638 | 2.4505 |
| 67.9 | 0.6317 | 0.1073 | 0.1882 | 0.2508 | 0.1839 | 0.2636 | 0.2274 | 0.0113 | 0.0135 |
| 69.3 | 3.0283 | 1.0725 | 0.8808 | 0.5264 | 0.6959 | 0.6076 | 0.8444 | 0.3076 | 0.4783 |
| 67.5 | 0.0520 | 0.0604 | 0.3166 | 0.3559 | 0.2036 | 0.3207 | 0.1524 | 0.0530 | 0.0163 |
| 68.7 | 0.8336 | 1.7577 | 1.5826 | 1.5683 | 2.2572 | 1.8026 | 2.3510 | 1.3764 | 1.7232 |
| 66.5 | 0.8947 | 0.7307 | 0.4007 | 0.6876 | 0.1449 | 0.3849 | 0.0859 | 0.1961 | 0.1252 |
| 67 | 1.0433 | 0.1030 | 0.4775 | 0.0259 | 0.0238 | 0.3505 | 0.0010 | 0.0884 | 0.0850 |
| 68.5 | 0.7418 | 0.3176 | 1.3479 | 0.5607 | 1.1094 | 1.2557 | 1.1238 | 1.6871 | 1.5715 |
| 69.3 | 1.2916 | 1.8698 | 0.8028 | 1.7506 | 1.7646 | 0.8330 | 1.4948 | 0.6877 | 0.5570 |
| Total | 40.4470 | 34.3289 | 31.3210 | 32.9453 | 33.0731 | 31.0651 | 32.7724 | 30.3520 | 30.4551 |
| MSE | 1.0932 | 0.9278 | 0.8465 | 0.8904 | 0.8939 | 0.8396 | 0.8857 | 0.8203 | 0.8231 |

From the above table, we get the error values of nine ARMA models i.e., ARMA (1, 1), ARMA (1, 2), ARMA (1, $3)$, ARMA $(2,1)$, ARMA $(2,2)$, ARMA $(2,3)$, ARMA (3, $1)$, ARMA $(3,2)$ and ARMA (3, 3). Computing mean square error values for nine ARMA models are 1.0931623 for ARMA (1, 1), 0.92780833 for ARMA (1, 2), 0.84651368 for ARMA $(1,3), 0.8904127$ for ARMA $(2,1)$, 0.8938671 for ARMA $(2,2), 0.8395974$ for ARMA $(2,3)$,
0.88574098 for ARMA (3, 1), 0.82032426 for ARMA (3, $2), 0.82311$ for ARMA (3, 3). An autoregressive moving average model $(3,2)$ possess lowest mean square error value of 0.82032426 . By using ARMA $(3,2)$ model, we estimating ' $\alpha$ ' value.

An ARMA $(3,2)$ model possesses lowest mean square error value of 0.8203 .

| $x_{i}$ | $\hat{x}_{i}$ | Error | abs error |
| :---: | :---: | :---: | :---: |
| 64.4 | 64.4846 | -0.0846 | 0.0846 |
| 66.2 | 64.5741 | 1.6259 | 1.6259 |
| 64.4 | 64.8073 | -0.4073 | 0.4073 |
| 63.9 | 64.5748 | -0.6748 | 0.6748 |
| 66.8 | 64.2647 | 2.5353 | 2.5353 |
| 65.5 | 65.1889 | 0.3111 | 0.3111 |
| 65.4 | 64.5224 | 0.8776 | 0.8776 |
| 63.3 | 64.2982 | -0.9982 | 0.9982 |
| 65.7 | 64.6453 | 1.0547 | 1.0547 |
| 65.1 | 65.2166 | -0.1166 | 0.1166 |
| 66.2 | 65.7104 | 0.4896 | 0.4896 |
| 65.5 | 65.2631 | 0.2369 | 0.2369 |
| 65.7 | 65.7068 | -0.0068 | 0.0068 |
| 64.5 | 65.1594 | -0.6594 | 0.6594 |
| 66.8 | 65.9376 | 0.8624 | 0.8624 |
| 66.1 | 65.8632 | 0.2368 | 0.2368 |
| 65.9 | 66.3059 | -0.4059 | 0.4059 |
| 65.4 | 65.4939 | -0.0939 | 0.0939 |
| 66.1 | 66.4533 | -0.3533 | 0.3533 |
| 66 | 66.0848 | -0.0848 | 0.0848 |
| 65.8 | 66.9652 | -1.1652 | 1.1652 |
| 64.5 | 66.3778 | -1.8778 | 1.8778 |
| 68.3 | 67.3678 | 0.9322 | 0.9322 |
| 68.6 | 67.509 | 1.091 | 1.091 |
| 67.4 | 67.5476 | -0.1476 | 0.1476 |
| 65.6 | 65.8788 | -0.2788 | 0.2788 |
| 65.3 | 66.5558 | -1.2558 | 1.2558 |
| 66 | 66.4854 | -0.4854 | 0.4854 |
| 66.5 | 68.1012 | -1.6012 | 1.6012 |
| 67.9 | 67.7935 | 0.1065 | 0.1065 |
| 69.3 | 68.7454 | 0.5546 | 0.5546 |
| 67.5 | 67.7303 | -0.2303 | 0.2303 |
| 68.7 | 67.5268 | 1.1732 | 1.1732 |
| 66.5 | 66.9428 | -0.4428 | 0.4428 |
| 67 | 67.2974 | -0.2974 | 0.2974 |
| 68.5 | 67.2011 | 1.2989 | 1.2989 |
| 69.3 | 68.4707 | 0.8293 | 0.8293 |
|  | Total | 2.5481 | 25.8839 |

The above table contains 4 columns, first column tells about original value, second column tells about estimated value of ARMA (3, 2), third column gives subtractions of estimated values from original values (errors) and fourth column says about modules of subtraction of estimated values from original values (absolute errors). $\alpha^{*}$ is estimated by using ratio of sum of errors of ARMA $(3,2)$ to sum of absolute errors of ARMA (3, 2).

$$
\begin{gathered}
\alpha^{*}=\frac{\sum\left(x_{i}-\hat{x}_{i}\right)}{\sum\left|x_{i}-\hat{x}_{i}\right|} \\
=0.9
\end{gathered}
$$

$$
\begin{aligned}
& \alpha^{*}+\beta^{*}=1 \\
& \quad \Rightarrow \beta^{*}=0.1
\end{aligned}
$$

Upon substituting estimated values of $\alpha^{*}, \beta^{*}$ in an equation, we get

$$
\begin{aligned}
& F_{t+1}=\alpha^{*} Y_{t}+\beta^{*} F_{t} \\
&=0.9 Y_{t}+0.1 S_{t}
\end{aligned}
$$

Simple exponential smoothing model is also fitted for data by taking $\alpha=0.1$ and $\beta=0.9$, is of the form

$$
\mathrm{F}_{\mathrm{t}+1}=0.1 \mathrm{Y}_{\mathrm{t}}+0.9 \mathrm{~F}_{\mathrm{t}}
$$

A mean square error criterion is used for obtaining which model is best between simple exponential smoothing and new simple exponential smoothing. For that purpose the following table explains forecasted values using exponential smoothing model \& new exponential smoothing model and their error squares are also discussed.

| Year | Annual | Exponential <br> Smoothing | $(\text { Error })^{2}$ | Exponential <br> Smoothing | $(\text { Error })^{2}$ |
| :--- | :--- | :--- | :--- | :--- | :--- |
| 1976 | 64.4 | 64.4000 | 0.0000 | 64.4000 | 0.0000 |
| 1977 | 66.2 | 64.5800 | 2.6244 | 66.0200 | 0.0324 |
| 1978 | 64.4 | 64.5620 | 0.0262 | 64.5620 | 0.0262 |
| 1979 | 63.9 | 64.4958 | 0.3550 | 63.9662 | 0.0044 |
| 1980 | 66.8 | 64.7262 | 4.3006 | 66.5166 | 0.0803 |
| 1981 | 65.5 | 64.8036 | 0.4850 | 65.6017 | 0.0103 |
| 1982 | 65.4 | 64.8632 | 0.2881 | 65.4202 | 0.0004 |
| 1983 | 63.3 | 64.7069 | 1.9794 | 63.5120 | 0.0450 |
| 1984 | 65.7 | 64.8062 | 0.7988 | 65.4812 | 0.0479 |
| 1985 | 65.1 | 64.8356 | 0.0699 | 65.1381 | 0.0015 |
| 1986 | 66.2 | 64.9720 | 1.5079 | 66.0938 | 0.0113 |
| 1987 | 65.5 | 65.0248 | 0.2258 | 65.5594 | 0.0035 |
| 1988 | 65.7 | 65.0924 | 0.3692 | 65.6859 | 0.0002 |
| 1989 | 64.5 | 65.0331 | 0.2842 | 64.6186 | 0.0141 |
| 1990 | 66.8 | 65.2098 | 2.5287 | 66.5819 | 0.0476 |
| 1991 | 66.1 | 65.2988 | 0.6419 | 66.1482 | 0.0023 |
| 1992 | 65.9 | 65.3589 | 0.2927 | 65.9248 | 0.0006 |
| 1993 | 65.4 | 65.3630 | 0.0014 | 65.4525 | 0.0028 |
| 1994 | 66.1 | 65.4367 | 0.4399 | 66.0353 | 0.0042 |
| 1995 | 66 | 65.4931 | 0.2570 | 66.0035 | 0.0000 |
| 1996 | 65.8 | 65.5238 | 0.0763 | 65.8204 | 0.0004 |
| 1997 | 64.5 | 65.4214 | 0.8490 | 64.6320 | 0.0174 |
| 1998 | 68.3 | 65.7092 | 6.7120 | 67.9332 | 0.1345 |
|  |  |  |  |  |  |


| 1999 | 68.6 | 65.9983 | 6.7687 | 68.5333 | 0.0044 |
| :--- | :--- | :--- | :--- | :--- | :--- |
| 2000 | 67.4 | 66.1385 | 1.5914 | 67.5133 | 0.0128 |
| 2001 | 65.6 | 66.0846 | 0.2349 | 65.7913 | 0.0366 |
| 2002 | 65.3 | 66.0062 | 0.4987 | 65.3491 | 0.0024 |
| 2003 | 66 | 66.0056 | 0.0000 | 65.9349 | 0.0042 |
| 2004 | 66.5 | 66.0550 | 0.1980 | 66.4435 | 0.0032 |
| 2005 | 67.9 | 66.2395 | 2.7573 | 67.7544 | 0.0212 |
| 2006 | 69.3 | 66.5456 | 7.5870 | 69.1454 | 0.0239 |
| 2007 | 67.5 | 66.6410 | 0.7379 | 67.6645 | 0.0271 |
| 2008 | 68.7 | 66.8469 | 3.4340 | 68.5965 | 0.0107 |
| 2009 | 66.5 | 66.8122 | 0.0975 | 66.7097 | 0.0440 |
| 2010 | 67 | 66.8310 | 0.0286 | 66.9710 | 0.0008 |
| 2011 | 68.5 | 66.9979 | 2.2563 | 68.3471 | 0.0234 |
| 2012 | 69.3 | 67.2281 | 4.2928 | 69.2047 | 0.0091 |
|  |  |  |  |  |  |
|  |  | Total | 55.5964 | Total | 0.7112 |
|  |  | MSE | 1.5026 | MSE | 0.0192 |

MSE of simple exponential smoothing is 1.5026 ; MSE of new simple exponential smoothing is 0.0192 ; Mean square error of new exponential smoothing model is less than the mean square error of simple exponential smoothing model. 0.0192 < 1.5026 . Therefore, we conclude that the new single
exponential smoothing model is the best model compared with general single exponential smoothing model and the parameter $\alpha=0.1$.These two models for the original annual data are shown in the figure-1.


## Figer-1

## SUMMARY AND CONCLUSIONS

We fitted nine ARMA models by changing autoregressive and moving average parameters. The mean square error of nine ARMA models are 1.0932 for ARMA (1, 1), 0.9278 for ARMA (1, 2), 0.8465 for ARMA (1, 3), 0.8904 for ARMA $(2,1), 0.8939$ for ARMA $(2,2), 0.8396$ for ARMA $(2,3)$, 0.8857 for ARMA $(3,1), \quad 0.8203$ for ARMA (3, 2), and 0.8231 for ARMA $(3,3)$. MSE of ARMA $(3,2)$ is smaller than all other ARMA models. So, we use ARMA (3, 2) model for estimating the parameter in simple exponential smoothing.
The fitted ARMA $(3,2)$ model is
$\mathrm{Y}_{\mathrm{t}}+0.131 \mathrm{Y}_{\mathrm{t}-1}-0.407 \mathrm{Y}_{\mathrm{t}-2}+0.461 \mathrm{Y}_{\mathrm{t}-3}-\varepsilon_{\mathrm{t}}-0.180 \varepsilon_{\mathrm{t}-1}+0.815$ $\varepsilon_{\mathrm{t}}-2+122.774=0$
$\alpha^{*}$ is estimated by the ratio of average of error to average of absolute error.

$$
\alpha^{*}=\frac{\text { mean error }}{\text { mean absoluteerror }}
$$

The estimated parameter $\alpha^{*}=0.9$
The forecast equation for simple exponential smoothing is

$$
\mathrm{F}_{\mathrm{t}+1}=0.1 \mathrm{Y}_{\mathrm{t}}+0.9 \mathrm{~F}_{\mathrm{t}}
$$

Fitted new exponential smoothing by using $\alpha^{*}$ is

$$
\mathrm{F}_{\mathrm{t}+1}=0.9 Y_{t}+0.1 \mathrm{~F}_{\mathrm{t}}
$$

A mean square error criterion is used for choosing best model between simple exponential smoothing and new simple exponential smoothing. MSE for simple exponential smoothing is 1.5026 and MSE for new exponential smoothing is 0.0192 . MSE of new exponential smoothing is less than MSE of simple exponential smoothing. So we conclude that the new simple exponential smoothing is better than the simple exponential smoothing model.

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## REFERENCES

[1] Cipra, T. (1992), "Robust exponential smoothing, Journal of Forecasting", 11, 57-69.
[2] Fried, R. (2004), "Robust filtering of time series with trends", Nonparametric statistics, 16,313-328.
[3] Holt, C.C. (1957), "Forecasting seasonal and trends by exponentially weighted moving averages", ONR research Memorandum 52, and R.J. (2004), International Journals of Forecasting, 20, 5-13.
[4] Spyros Makridakis, Steven C. Wheelwright and Rob J. Hyndman, (1998) "Forecasting methods and applications", Third Edition, Wiley india Pvt. Ltd., New Delhi.
[5] Winters, P.R. (1960), "Forecasting sales by exponentially weighted moving averages", Management Science, 6, 324-342.
[6] Box, G.E.P. and G.M. Jenkins (1970) time series analysis: Forecasting and control, San Francisco: Holden-Day.
[7] Box, G.E.P. and D.A. Pierce (1970) Distribution of the residual autocorrelations in autoregressiveintegrated moving-average time series models, Journal of the American Statistical Association, 65, 1509-1526.
[8] Mc Kenzie, E. (1984) General exponential smoothing and the equivalent ARIMA process, Journal of Forecasting, 3, 333-334.
[9] Mc Kenzie, E. (1986) Error analysis for Winters’ additive seasonal forecasting system, International Journal of Forecasting, 2, 373-382.
[10]http://www.srh.noaa.gov/graphicast.php?site=fwd

